

# Optimal Control Variates for MCMC

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## Objective:

Minimize asymptotic variance of MCMC estimates using control variates

## Introduction

In many applications, we need to evaluate expectations that are difficult to compute:

$$\eta = \int c(x)\pi(x)dx$$

$c: \mathbb{R}^\ell \rightarrow \mathbb{R}$  is a measurable function,  $\pi$  is a probability density in  $\mathbb{R}^\ell$ .

**Markov-Chain Monte Carlo** (MCMC) methods provide numerical algorithms to obtain empirical estimates:

$$\eta(t) = \frac{1}{t} \int_0^t c(\Phi(s)) ds$$

$\Phi$  is a Markov process with steady state distribution  $\pi$ .

Rate of convergence captured by Central Limit Theorem, with asymptotic variance

$$\gamma^2 = \lim_{t \rightarrow \infty} \mathbb{E} \left[ \left( \frac{1}{\sqrt{t}} \int_0^t (c(\Phi(s)) - \eta) ds \right)^2 \right]$$

## Langevin Diffusion

**Langevin diffusion** is an  $\ell$ -dimensional gradient flow with “noise”:

$$d\Phi(t) = -\nabla \mathcal{U}(\Phi(t)) dt + \sqrt{2} dW(t)$$

$W = \{W(t) : t \geq 0\}$  is a standard Brownian motion on  $\mathbb{R}^\ell$

Under general conditions, this diffusion is reversible, with unique invariant density,  $\pi = e^{-\mathcal{U} + \Lambda}$ , where  $\Lambda$  is a normalizing constant.

Differential generator of the Langevin diffusion,

$$\mathcal{D}f := -\nabla \mathcal{U} \cdot f + \Delta f, \quad f \in C^2$$

Poisson’s equation for the Langevin diffusion:

$$\mathcal{D}h = -\tilde{c}, \quad \tilde{c} = c - \eta$$

Representation for  $\gamma^2$  in terms of the solution to Poisson’s equation  $h$ :

$$\gamma^2 = 2\langle h, \tilde{c} \rangle_\pi$$

For Langevin diffusion, on integration by parts, we have

$$\gamma^2 = 2\langle h, \tilde{c} \rangle_\pi = -2\langle h, \mathcal{D}h \rangle_\pi = 2\langle \nabla h, \nabla h \rangle_\pi = 2\|\nabla h\|_\pi^2$$

## Control variates and TD-learning

A **control variate**  $n$  has zero mean with respect to the density  $\pi$ , i.e.  $\int n(x)\pi(x)dx = 0$ . Eg:  $n = \mathcal{D}f$ , for a “nice”  $f$ , i.e.  $f \in L_1(\pi)$  and  $f \in C^2$ .

We look at an improved estimator for  $\eta$ ,

$$\eta_\theta(t) := \frac{1}{t} \int_0^t c_\theta(\Phi(s)) ds, \quad \text{where } c_\theta := c + \underbrace{\mathcal{D}h_\theta}_{\text{control variate}}$$

$$h_\theta := \sum_{i=1}^d \theta_i \psi_i, \quad \text{where } \theta \in \mathbb{R}^d \text{ and } \psi_i \text{ is the } i^{\text{th}} \text{ basis function.}$$

Ideally, we prefer  $\mathcal{D}h$ , but  $h$  is difficult to solve. Approximating  $h$  is hard and is the subject of TD-learning:

$$\min_{\theta} \|h - h_\theta\|_\pi^2$$

Our goal is to find the “optimal” control variate that minimizes asymptotic variance, i.e.

$$\min_{\theta} \gamma_\theta^2 := \min_{\theta} \|\nabla h - \nabla h_\theta\|_\pi^2$$

which is just as hard in general. However for Langevin, we have a simple solution.

## Algorithm

$$\theta^* := M^{-1}b, \quad \theta(t) = M^{-1}(t)b(t)$$

$$M_{ij} := \sum_{k=1}^d \langle \partial_k \psi_i, \partial_k \psi_j \rangle_\pi, \quad M(t) = M(0) + \int_0^t \nabla \psi(\Phi(s)) \nabla \psi^T(\Phi(s)) ds,$$

$$b_i := \langle \psi_i, \tilde{c} \rangle_\pi. \quad b(t) = \int_0^t \psi(\Phi(s)) \tilde{c}(\Phi(s)) ds.$$

Exact solution Monte Carlo approximation

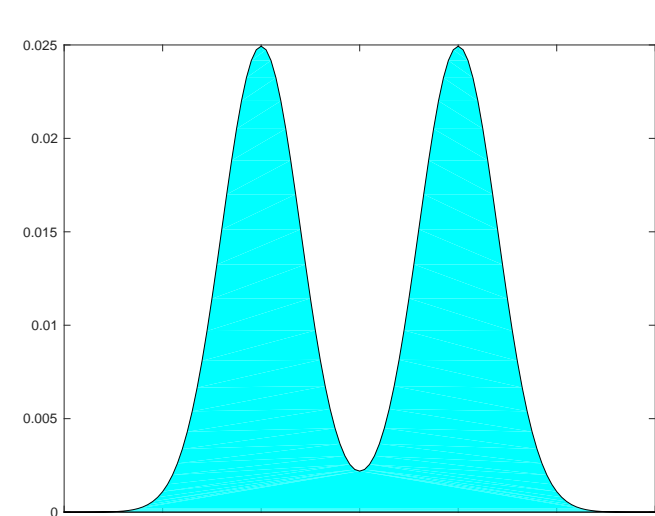
$$\lim_{t \rightarrow \infty} \theta(t) = \theta^* \quad \text{with probability one.}$$

## Numerical Example

**Target density**  $\pi$  is a mixture of Gaussians

$$\pi = \sum_{m=1}^M w_m p_m \quad \text{where } \sum_{m=1}^M w_m = 1, \quad p_m := \mathcal{N}(\mu_m, \sigma_m^2)$$

Eg:  $M = 2 \quad \mu_1 = -1, \mu_2 = 1 \quad \sigma_1 = \sigma_2 = 0.4 \quad w_1 = w_2 = 0.5$



**Parameterization:** A linear parameterization of the form

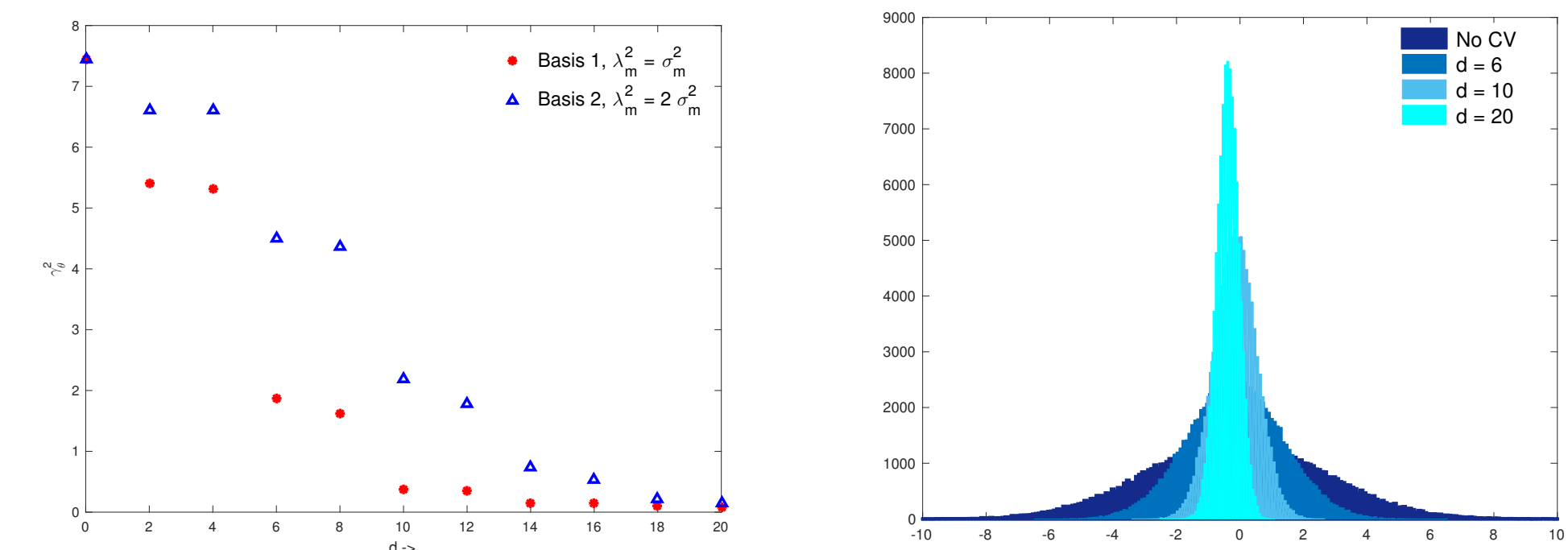
$$h_\theta(x) = \sum_{i=1}^d \theta_i \psi_i(x)$$

$$\{\psi_i(x) : x \in \mathbb{R}, 1 \leq i \leq d\} = \{x^k q_m(x) : 0 \leq k \leq d_0, m = 1, 2, d = 2d_0\}$$

$$q_m \sim \mathcal{N}(\mu_m, \lambda_m^2) \quad \text{such that } \lambda_m^2 \geq \sigma_m^2.$$

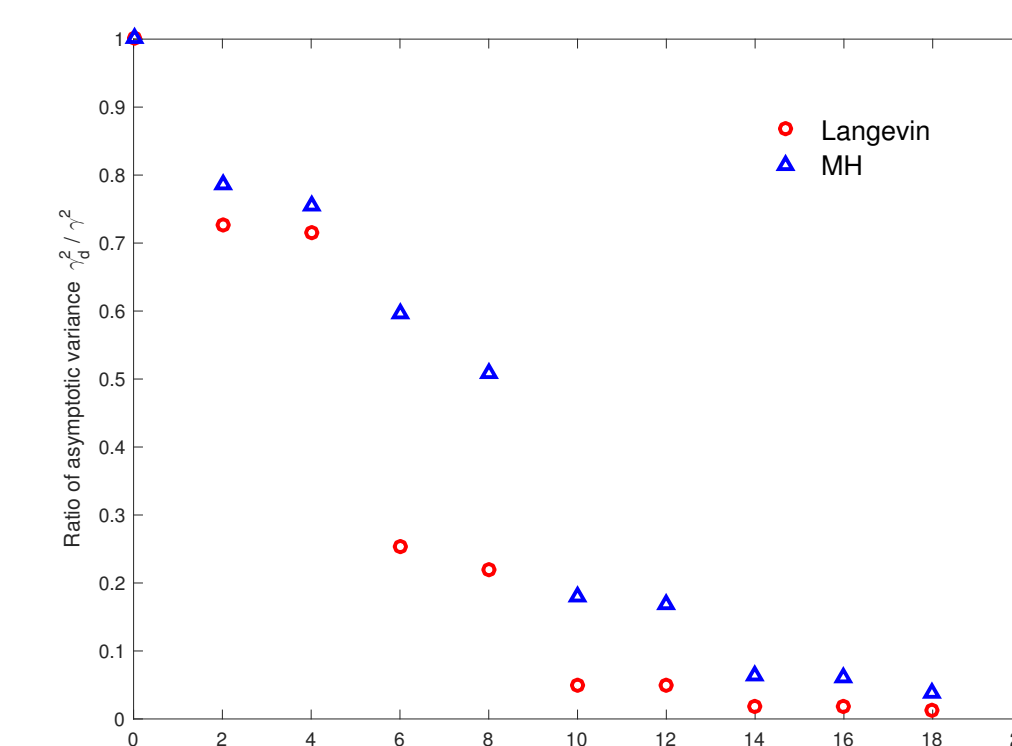
## Asymptotic variance reduction

**Asymptotic variance reduction with  $d$ :** Basis 1 with  $\lambda_m^2 = \sigma_m^2$  V Basis 2 with  $\lambda_m^2 = 2\sigma_m^2$ .



## Application to Metropolis-Hastings algorithm

$\mathcal{D}h_\theta$  remains a control variate for the MH algorithm with target density  $\pi$ . Although not “optimal”, reduction in asymptotic variance achieved is similar to Langevin case.



## Variance v Asymptotic variance

In prior research [1, 2] the control variate is chosen by minimizing the ordinary variance rather than the asymptotic variance.

Well-motivated only if the samples are i.i.d. The unbiased estimator is based on a function  $c_\theta$ :

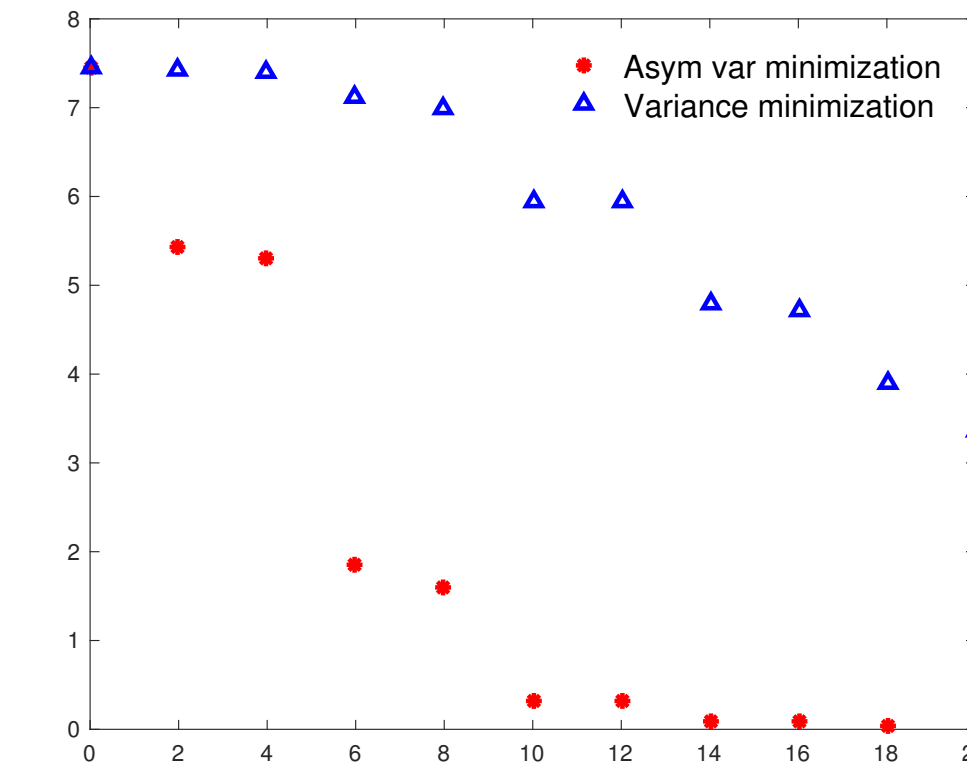
$$c_\theta = c + \mathcal{D}h_\theta, \quad \text{where } h_\theta = \sum_{i=1}^d \theta_i \psi_i = \vartheta^T \psi$$

Optimal parameter  $\vartheta^*$  is obtained by minimizing the variance:

$$\vartheta^* := \arg \min_{\vartheta} \sigma_{c_\theta}^2 = \arg \min_{\vartheta} \|c_\theta - \eta\|_\pi^2 = -\overline{M}^{-1} \overline{b}$$

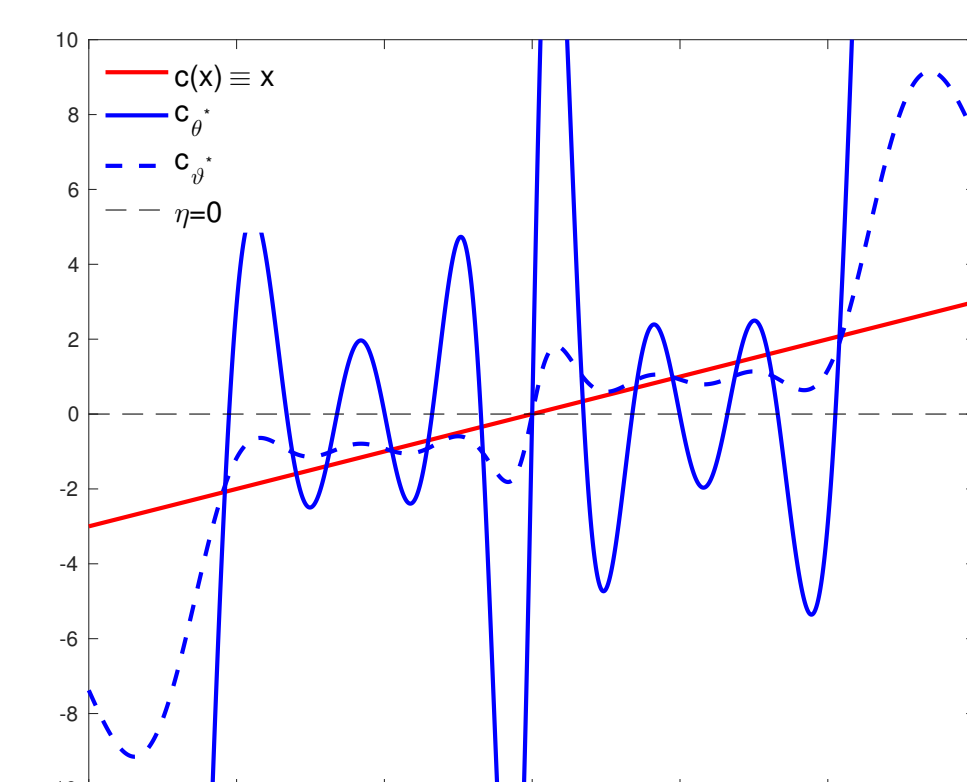
$$\overline{M} := \langle \mathcal{D}\psi, \mathcal{D}\psi \rangle_\pi, \quad \overline{b} := \langle \mathcal{D}\psi, \tilde{c} \rangle_\pi$$

But, **variance minimization  $\neq$  asymptotic variance minimization.**



Control variates  $c_\theta^*$  and  $c_\vartheta^*$  imply that “**oscillations are good**”.

$$\int_{-\infty}^0 c_{\theta^*}(x)p(x) dx \approx \int_0^{\infty} c_{\vartheta^*}(x)p(x) dx \approx \eta = 0$$



## Conclusion

Optimization of asymptotic variance is remarkably simple for the Langevin diffusion.

In parallel research, these techniques are applied to gain approximation for the feedback particle filter.

Ongoing research includes:

- Basis selection for multidimensional models.
- Investigating a basis independent approach using reproducing kernels.

## References

- [1] Antonietta Mira, Reza Solgi, and Daniele Imparato. Zero variance markov chain monte carlo for bayesian estimators. *Statistics and Computing*, 23(5):653–662, September 2013.
- [2] Chris J. Oates, Mark Girolami, and Nicolas Chopin. Control functionals for monte carlo integration. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 79(3):695–718, 2017.

## Acknowledgements